

Problem of the Week

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Problem. Let $m_0 = n$. For each $i \geq 0$, choose a number x_i in $\{1, \dots, m_i\}$ uniformly at random and let $m_{i+1} = m_i - x_i$. This gives a random vector $\mathbf{x} = (x_0, x_1, \dots)$. For each $1 \leq k \leq n$, let X_k be the number of occurrences of k in the vector \mathbf{x} . For each $1 \leq k \leq n$, let Y_k be the number of cycles of length k in a permutation of $\{1, \dots, n\}$ chosen uniformly at random. Prove that X_k and Y_k have the same distribution.

Proof. For each permutation $\pi \in S_n$, order cycle depending on order of its least element.

Define $c(\pi) := \mathbf{x} = \{x_0, x_1, \dots, x_{l-1}\}$ where x_k is length of k -th cycle.

For any vector $\mathbf{x} = \{x_0, \dots, x_{l-1}\}$, define $m_i(\mathbf{x}) := \sum_{k=i}^{l-1} x_k$ and $m_l(\mathbf{x}) = 0$, $t(\mathbf{x}) = \{x_1, \dots, x_{l-1}\}$.

Clearly, $m_i(\mathbf{x}) - m_{i+1}(\mathbf{x}) = x_i$ and $m_{i+1}(\mathbf{x}) = m_i(t(\mathbf{x}))$. By definition of $c(\pi)$, $m_0(c(\pi)) = n$.

Claim. Probability of obtaining $\mathbf{x} = c(\pi)$ is $P(\mathbf{x}) = \prod_{i=0}^{l-1} \frac{1}{m_i(\mathbf{x})}$.

Enough to show there are $n! P(\mathbf{x})$ permutations with $\mathbf{x} = c(\pi)$, since the probability of choosing specific permutation π is identically $1/n!$.

Apply induction on l . By definition, 0^{th} cycle always contains 1. Choose remaining $x_0 - 1$ elements in

the first cycle in order, which gives ${}_{n-1}P_{x_0-1} = \frac{(n-1)!}{(n-x_0)!}$ cases.

If $m = 1$, done ($\because x_0 = n$, ${}_{n-1}P_{n-1} = (n-1)! = \frac{n!}{n} = \frac{n!}{m_0(\mathbf{x})} = n! P(\mathbf{x})$).

If $m > 1$, re-number the remaining of $\{1, \dots, n\}$ excluding 0^{th} cycle, without changing order, so that it becomes $\{1, \dots, n-x_0\}$. As we are working with permutation, this does not affect number of cases.

By induction hypothesis, number of cases of choosing the other cycles is $(n-x_0)! P(t(\mathbf{x}))$.

Multiplying number of choices of first cycle, we obtain total of $(n-1)! P(t(\mathbf{x}))$ possibilities.

Then, $(n-1)! P(t(\mathbf{x})) = n! \times \frac{P(t(\mathbf{x}))}{m_0(\mathbf{x})} = n! P(\mathbf{x})$. \square

Now let $M(\mathbf{x}) = \{m_0(\mathbf{x}), m_1(\mathbf{x}), \dots, m_l(\mathbf{x})\}$, which is upper bound of x_i in the process of X_k .

Claim. Probability of obtaining \mathbf{x} from process of X_k is $P(\mathbf{x})$.

Consider n to be $m_0(\mathbf{x})$. Apply induction on l .

For $l = 1$, n has to be chosen first from $\{1, \dots, n\}$ and it's done. Probability is $1/n = P(\mathbf{x})$.

For $l > 1$, x_0 has to be chosen first from $\{1, \dots, n\}$, which gives $1/n$. After that, it has to choose $t(\mathbf{x})$.

Since $m_0(t(\mathbf{x})) = m_1(\mathbf{x}) = m_0 - x_0$, this is identical to choosing $t(\mathbf{x})$ from scratch.

By induction hypothesis, choosing $t(\mathbf{x})$ has probability of $P(t(\mathbf{x}))$. In total, $\frac{1}{n} \times P(t(\mathbf{x})) = P(\mathbf{x})$. \square

Observe that for each vector \mathbf{x} with $m_0(\mathbf{x}) = n$ and $c(\pi) = \mathbf{x}$, both X_k and Y_k are occurrences of k in \mathbf{x} , since the number of k -cycle in π can be counted as occurrence of k in $c(\pi)$.

Obviously, $c(\pi)$ for $\pi \in S_n$ spans entire $\{\mathbf{x} : x_k > 0 \text{ for } k=1, 2, \dots, m, m_0(\mathbf{x}) = n\}$, i.e. both have the same range of \mathbf{x} . Because $X_k = Y_k$ given $c(\pi) = \mathbf{x}$, and probability of choosing $c(\pi) = \mathbf{x}$ for both process is the same, X_k and Y_k should have the same distribution.